



Faking Effects on the Factor Structure of a Quasi-Ipsative Forced-Choice Personality Inventory

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ABSTRACT

Research has shown that faking behavior affects the factor structure of single-stimulus (SS) personality measures. However, no published research has analyzed the effects of this phenomenon on the factor structure of forced-choice (FC) personality inventories. This study examines the effects of faking, induced in a laboratory setting, on the construct validity of a quasi-ipsative FC personality inventory based on the Five-Factor Model. It also examines the moderator effect of the type of experimental design (between-subject and within-subject design) on factor analyses. The results showed that (a) data fit to a structure of five-factors in the two conditions (honest and faking) in both experimental designs; (b) model fit indices are also good or excellent in all cases; and (c) Burt-Tucker's congruence coefficients between convergent factors of conditions analyzed are very high. These findings provide evidence that the quasi-ipsative FC format is a robust instrument that controls the effects of faking on factor structure. Finally, we discuss theoretical and practical implications of these findings for personnel selection and assessment.

Los efectos del *faking* en la estructura factorial de un inventario de personalidad de elección forzosa quasi-ipsativo

RESUMEN

La investigación ha demostrado que el *faking* o *falseamiento* afecta a la estructura factorial de las medidas de personalidad de estímulo único (single stimulus; SS). Sin embargo hasta la fecha no se ha publicado ninguna investigación que haya analizado los efectos de este fenómeno en la estructura factorial de los cuestionarios de personalidad de elección forzosa (forced-choice, FC). Este estudio examina los efectos del falseamiento, inducido en un entorno de laboratorio, en la validez de constructo de un cuestionario de personalidad de elección forzosa quasi-ipsativo basado en el modelo de los cinco grandes factores de personalidad. También examina el efecto moderador del tipo de diseño experimental (inter-sujeto e intra-sujeto) en los análisis factoriales. Los resultados mostraron que (a) los datos se ajustan a una estructura de cinco factores en las dos condiciones (honesta y *faking*) en ambos diseños experimentales, (b) los índices de ajuste del modelo son buenos o excelentes en todos los casos y (c) los coeficientes de congruencia de Burt-Tucker entre los factores convergentes de las condiciones analizadas son muy altos. Estos hallazgos proporcionan evidencia de que el formato de elección forzosa quasi-ipsativo es un instrumento robusto que controla los efectos del *faking* en la estructura factorial. Finalmente se analizan las implicaciones teóricas y prácticas de estos resultados en la selección y evaluación de personal.

Personality inventories are an assessment procedure widely used in organizational settings, especially for high-stakes selection decisions. (e.g., Alonso et al., 2015; García-Izquierdo et al., 2019; García-Izquierdo et al., 2020; Golubovich et al., 2020; Heller, 2005; Rothstein & Goffin, 2006; Sackett, 2017). Empirical evidence has shown that personality instruments based on the Five-Factor model are valid predictors of relevant organizational and academic criteria, such as overall job performance, job satisfaction, leadership, and counterproductive behavior, among others (see for instance, Barrick

et al., 2001; Bartram, 2005; Cuadrado et al., 2020, 2021; Delgado-Rodríguez, 2018; Judge et al., 2013; Lado & Alonso, 2017; Poropat, 2009; Salgado, 1997, 2003; Salgado et al., 2015; Salgado et al., 2013; Salgado et al., 2020). Conscientiousness and emotional stability predicted all the criteria analyzed, while the other three factors (extraversion, openness to experience, and agreeableness) predicted some specific criteria in occupational categories. For instance, extraversion is a predictor of performance in jobs that require social interaction, agreeableness is a performance predictor for occupations

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oriented to cooperation and to help others, and openness to experience is a relevant predictor of performance for jobs requiring, for example, high levels of creativity.

Typically, personality has been assessed with single-stimulus (SS) questionnaires. SS instruments are characterized by the fact that individuals must rate every item of the test indicating their level of agreement with it (through Likert scales, true/false, or yes/no) in order to describe their personality. However, due to the characteristics of this type of answer, SS questionnaires have frequently been criticised for their potential susceptibility to faking behavior (e.g., Christiansen et al., 2005; Griffith & McDaniel, 2006; McFarland & Ryan, 2000; Rosse et al., 1998; Viswesvaran & Ones, 1999). Meta-analytic research of Viswesvaran and Ones (1999), Birkeland et al. (2006), and Salgado (2016) pointed out that individuals can distort their scores on SS instruments if they are motivated to fake.

In order to reduce or control the faking effects on personality assessment, forced choice (FC) personality inventories have been suggested as an alternative. Meta-analytic evidence revealed that when the Big Five personality dimensions are assessed with one specific class of FC personality questionnaires, i.e., quasi-ipsative FC personality questionnaires, they showed larger validity coefficients for predicting job performance than did SS personality inventories (e.g., Salgado, 2017; Salgado & Tauriz, 2014; Salgado et al., 2015). Recent meta-analytic studies have also shown that quasi-ipsative FC personality questionnaires are robust against faking effects (Cao & Drasgow, 2019; Martínez, 2019; Salgado & Lado, 2018).

However, some issues concerning quasi-ipsative FC personality inventories remain unexamined, for instance, whether the factor structure (construct validity) of FC personality inventories is resistant to the effects of faking. This article reports a study that examines the effects of faking on the factor structure of a quasi-ipsative FC personality inventory. To the best of our knowledge, no previous research has been undertaken to analyze this issue. Specifically, this study examines whether the quasi-ipsative FC inventory presents measurement invariance between honest and faking response-conditions. In the second place, the study examines whether the type of experimental design (within-subject and between-subject design) is a moderator of the effects of faking on the construct validity of quasi-ipsative FC.

Faking Behavior

McFarland and Ryan (2000) defined faking as “a type of response bias where an individual consciously distorts answers to be viewed favorably” (p. 813). Along the same lines, Paulhus (2002) described this behavior as the intention to show an “overly positive self-description” (p. 50) and, more recently, Ziegler et al. (2012) provided a more comprehensive definition of faking, describing this behavior as “a set of responses aimed at providing a portrayal of the self that helps a person to achieve personal goals. Faking occurs when this response set is activated by situational demands and personal characteristics to produce systematic differences in test scores that are not due to the attribute of interest” (p. 8). Hence, from these definitions, faking can be understood as intentional behavior to distort responses, by choosing the most socially desirable answer, to offer an image that favors the individual in the evaluation process.

Due to the adverse characteristics of this behavior, there is considerable concern to understand the potential negative effects that faking could have on SS personality questionnaires (see, for instance, Birkeland et al., 2006; Salgado, 2016; Viswesvaran & Ones, 1999). Specifically, research has focused on the effects of faking behavior on (a) the scores, (b) the reliability, (c) the validity, and (d) the ranking of candidates (hiring decisions).

The effects on the mean of scores are among the most studied. The meta-analyses of Viswesvaran and Ones (1999), Birkeland et al.

(2006), Hooper (2007), and Salgado (2016) have shown that faking causes an increase in the scores of personality measures and that this effect is greatest for conscientiousness and emotional stability in all cases. Likewise, meta-analytic evidence showed that faking also reduces the magnitudes of standard deviations of scores obtained with SS questionnaires (Salgado, 2016; Viswesvaran & Ones, 1999). These two simultaneous effects are caused by an artificial homogenization of samples, producing on average a reduction of the range of scores obtained by the individuals. Therefore, it would be more difficult to differentiate people's suitability to be hired (Salgado, 2016).

In addition, empirical evidence has shown that reliability of SS questionnaires could be also affected by faking. The findings on this issue have revealed that reliability decreases significantly, that is, the degree of error contained in the measure increases when faking occurs (see, for instance, Salgado, 2016; Stark et al., 2012; Van Iddekinge et al., 2005). Consequently, the scores obtained with SS measures could be less reliable if faking is affecting data because measurement standard error is larger.

Regarding criterion validity, findings of the recent meta-analysis by Salgado (2016) has shown that faking causes attenuation of criterion validity in SS questionnaires. When individuals commit faking, a more imperfect measurement of the personality variables is obtained, which necessarily implies a reduction in the predictive power of these instruments.

Finally, in relation to construct validity, the effects found were less consistent and were examined for SS questionnaires only. Some studies have shown that faking does not affect factor structure of personality measures (Ellingson et al., 2001; Marshall et al., 2005; Michaelis & Eysenck, 1971; Smith & Ellingson, 2002; Smith et al., 2001), while other primary studies have indicated that faking behavior modifies factor structure. Some of these studies found that faking led to a decrease in the number of factors (see Ellingson et al., 1999; Pauls & Crost, 2005; Van Iddekinge et al., 2005), whereas other studies found that faking produced additional factors (Cellar et al., 1996; Schmit & Ryan, 1993). The psychometric theory of faking effects (Salgado, 2016) proposed that faking can produce both effects depending on whether faking affects a single scale (univariate range restriction) or several scales (multivariate range restriction). In the first case, faking will lead to smaller loadings in factor structure. In the second case, faking will produce additional factors. In both cases, effects of faking produce a lack of equivalence between the two conditions (honest and faking) in SS personality inventories.

In summary, the empirical evidence has clearly shown that faking is a source of variance that affects the psychometric properties of SS personality measures. Consequently, this phenomenon has a negative impact on the hiring-decisions in selection processes because individuals who better distort their answers (inflating their scores) get higher positions in the selection ranking, relegating the subjects who have answered honestly to lower positions. However, this evidence refers only to SS measures and does not necessarily apply to quasi-ipsative FC personality inventories. Therefore, it seems appropriate to examine the effects on the factor structure of quasi-ipsative FC personality inventories.

Forced-Choice Inventories

Forced-choice (FC) personality inventories are personality measures designed to reduce faking behavior and to be an alternative to traditional SS instruments (see, for instance, Baron, 1996; Bartram, 1996; Borislow, 1958). The first publications about FC inventories dates back to the 1940s (see Gordon, 1951; Hicks, 1970; Zavala, 1965); nonetheless, the interest in examining the resistance of these measures to the effects of faking has increased over the last decades. The main characteristic of FC inventories is that each item of the questionnaire presents various alternatives (usually pairs, triads,

or tetrads) with a similar degree of social desirability from which individuals must choose the alternative that describes them best and, in some cases, the alternative that least describes them. As all the answer options are socially attractive, it is more difficult for the applicants to distort the results. Consequently, their responses reflect the real choices that an individual would make. Therefore, the use of FC personality inventories should reduce the effects of faking (see Brown & Maydeu-Olivares, 2013; Christiansen et al., 2005; Converse, et al. 2006; Dilchert & Ones, 2012; Jackson et al., 2000).

Depending on how the answer-choice is made, FC personality inventories can provide three types of scores (i.e., normative, ipsative, and quasi-ipsative or partially ipsative scores), each of them with specific psychometric properties (Clemans, 1966; Hicks, 1970; Meade, 2004). FC inventories that provide normative scores are characterized by presenting only answers-alternatives of the same dimension in each item, that is, each item evaluates just one personality factor and the same alternatives are never used to represent different factors. In the case of ipsative scores, the individual must assess all the alternatives given for each item. Therefore, the score for each dimension depends on an individual's scores on the other graded dimensions and, consequently, the sum of the scores obtained for each individual is a constant. Finally, quasi-ipsative scores include measures that do not meet all criteria to be purely ipsative but present some characteristics associated with them (Clemans, 1966). Specifically, Hicks (1970) indicated that a quasi-ipsative score is obtained when some of the following criteria apply: (1) individuals only partially order the alternatives; (2) scales have different numbers of items; (3) not all of the items ranked by respondents are scored; (4) scales are scored differently for differing respondent characteristics; (5) items differ in how they are weighted; (6) some ipsative scales are deleted when data are analyzed; and (7) the inventory includes normative sections.

Additionally, two types of quasi-ipsative FC inventories can be distinguished, depending on whether they provide algebraically dependent or non-algebraically dependent scores. In the first type, there is a metric dependence between personality scales and, therefore, there is some degree of ipsativization of scores. In the second type, the score for each personality factor is algebraically independent of the score for the other factors.

Meta-analytical evidence has shown that the three types of FC personality inventories are valid predictors of occupational and academic criteria, obtaining similar or higher effect sizes than those produced with SS measures (Bartram, 2005, 2007; Salgado, 2017; Salgado et al., 2015; Salgado & Táuriz, 2014). In particular, quasi-ipsative FC inventories stand out above SS inventories and ipsative and normative FC inventories (Lee et al., 2018; Salgado, 2017; Salgado et al., 2015; Salgado & Táuriz, 2014) even in faking response-conditions (Martínez, 2019).

Forced-choice Inventories and Effects of Faking

Empirical evidence produced about the effectiveness of FC inventories to control the effects of faking suggests that FC personality formats are valid instruments for reducing response distortion. So far, the overall results of meta-analyses by Adair (2014), Cao and Drasgow (2019), Nguyen and McDaniel (2000), and Martínez (2019) support the conclusion that FC inventories show resistance to the effects of faking in experimental and occupational settings. Although the three types of FC show resistance to faking, the FC quasi-ipsative answer format proved to be more robust against this phenomenon than ipsative and normative FC inventories (Martínez, 2019; Salgado & Lado, 2018).

Regarding the effects of faking on construct validity of FC personality inventories, this issue has not been examined in previous research. Instead, extensive research has focused on examining convergent validity (correlations between scales) between SS personality

inventories and FC personality inventories under honest conditions. Findings showed evidence of the strong equivalence between both formats, pointing out that SS and FC instruments measure essentially the same personality constructs (Brown & Maydeu-Olivares, 2011, 2013; Cattell & Brennan, 1994; Joubert et al., 2015; Lee et al., 2019; Lee et al., 2018; Morillo et al., 2019; Zhang et al., 2020). Recently, Otero et al. (2020) simultaneously examined the factor structure of an SS personality questionnaire (IP/5F; Salgado, 1998) and a quasi-ipsative FC questionnaire (QI5F_tri; Salgado, 2014). The results confirmed a five-factor structure in both cases and demonstrated the validity of these inventories to assess Big Five factors.

Therefore, it becomes necessary to investigate the effects that faking can have on the factor structure of quasi-ipsative FC inventories, since previous research has revealed the suitability of this specific FC format, when compared to the other FC types, for controlling the effects of faking on scores even in real-life personnel selection settings (Cao & Drasgow, 2019; Martínez, 2019; Salgado & Lado, 2018).

Aims of the Study and Research Hypotheses

The current study aims to contribute to the knowledge of the effects of faking on the construct validity of quasi-ipsative FC inventories. Specifically, this study has two main objectives. First, to analyze the effects of faking on the factor structure of a non-dependent algebraically quasi-ipsative FC inventory; in other words, to analyze the invariance or equivalence of the measure under honest and faking conditions (Millsap, 2011). Secondly, to examine whether the experimental design (within- and between-subject designs) acts as a moderator of the magnitude of the effects of faking on the factor structure of this format.

As noted above, faking can reduce the number of factors or produce additional ones in SS measures (Salgado, 2016), depending on whether faking is univariate or multivariate. Although this issue has not been examined in FC personality measures, quasi-ipsative FC inventories have been shown to be more resistant to the effects of faking on scores and predictive validity SS measures (Martínez, 2019; Otero et al., 2020; Salgado et al., 2015; Salgado & Táuriz, 2014). Therefore, quasi-ipsative FC inventories are likely to also show resistance to the effects of faking on factor structure. Consequently, we propose the following hypothesis:

Hypothesis 1: Quasi-ipsative FC shows equivalent factor structure in honest and faking response-conditions.

Likewise, previous studies have indicated that the experimental design of the study (within-subject and between-subject designs) affects the magnitude of the effects of faking on the scores of SS questionnaires and FC formats, but these differences were smaller in effect sizes for quasi-ipsative FC inventories (Martínez, 2019; Viswesvaran & Ones, 1999). Based on the previous empirical evidence, we posit the following hypothesis:

Hypothesis 2: Quasi-ipsative FC inventories show equivalent factor structure in the within-subject design and in the between-subject design under honest and faking response-conditions.

Method

Sample and Experimental Design

Participants were 1,141 students from the University of Santiago de Compostela. The average age was 21.05 years ($SD = 4.04$) and 67.13% were women (766 participants). Small group sessions were organized to respond to the inventory. Participation was voluntary and all subjects provided informed consent to participate in the study.

Regarding the experimental design, 43% ($n = 490$) of the subjects participated in a within-subject design, in which all participants

completed the questionnaire honestly and under conditions that induced them to commit faking. The remaining 57% participated in a design of two independent groups (between-subject design) in which 449 subjects completed the questionnaire in the honest condition and 202 participants under faking instructions.

Measures

Q15F_tri. Personality was evaluated with the quasi-ipsative FC questionnaire, Q15F_tri, developed by [Salgado \(2014\)](#). This questionnaire consists of 140 triads that evaluate the Big Five personality factors. Each of the Big Five is evaluated using 28 items and each item contains three alternatives. Each alternative can be a short sentence or an adjective. All the alternatives are similar on social desirability and are presented in positive form, that is, there are no negative alternatives. No item presents two phrases or adjectives that evaluate the same personality factor, and no item is used to measure two or more personality factors. For example, an item may include a phrase that evaluates openness to experience, another that evaluates emotional stability, and another that evaluates agreeableness. The Q15F_tri implements [Horn's \(1971\)](#) strategy of quasi-ipsatization, so that the items used to evaluate a personality dimension are not used to evaluate other personality dimensions. Thus, the score for each of the Big Five is algebraically independent of the score for the other personality factors even though the format score is quasi-ipsative. In each item, individuals must choose the alternative that best describes them and the alternative that least describes them. An example of an item from this questionnaire is the following: "I am a person (a) who is open-minded; (b) who is a perfectionist; (c) who does not usually lose their temper."

The technical manual of Q15F_tri reports that internal consistency coefficients for emotional stability (ES), extraversion (EX), openness to experience (OE), agreeableness (A), and conscientiousness (C) were .71, .73, .80, .66, and .80, respectively. The technical manual also presents test-retest reliabilities for a four-week interval as .91, .90, .79, .65, and .72 for ES, EX, OE, A, and C, respectively. [Otero et al. \(2020\)](#) also reported convergent-discriminant validity evidence using a SS personality inventory.

Procedure

The quasi-ipsative FC inventory was used under two experimental response-conditions: honest and faking. In the honest condition, participants followed the instructions that are described below:

In the following questionnaire you will be presented with sets of phrases grouped into triads. Try to rank them by first identifying the one that best describes you, the one that second best describes you, and finally the one that describes you least. In each item, mark a plus sign (+) next to the phrase that best describes you and a minus sign (-) next to the phrase that least describes you. You should leave blank the one you considered second.

For the faking condition, test instructions were slightly modified in such a way that participants were encouraged to fake. The instructions were as follows:

Next, you will be presented with sets of phrases grouped into triads. Try to rank them by first identifying the one that best describes you, the one that second best describes you, and finally the one that describes you least. When answering, imagine that you are in the last step of a selection process for a very attractive job. Since it offers you a great opportunity to advance your professional career, you want to get that job. To do this, you must answer this test trying to give the best image of yourself to get that job. In each item, mark a plus sign (+) next to the alternative that best describes you and a minus sign (-) next to the alternative that

least describes you. You should leave blank the one you considered second best.

In both conditions, the questionnaire was administered (1) in paper-and-pencil format and (2) in computer format (using the Inquisit program; [Millisecond, 2016](#)). In both cases, instructions were the same regardless of the administration format that was used, and the two formats were not mixed in the same small group session. In addition, participants only had access to the questionnaire during the time they attended the study and they responded using only one of the administration formats.

Statistical Analyses

For the statistical analyses, the 140 items that make up the Q15F_tri questionnaire were grouped arbitrarily (it was not done randomly in the strictest statistical sense) into 20 clusters or compounds. Each cluster included 7 items from the instrument that evaluated the same personality factor. Thus, the 28 items that evaluated each personality factor were divided into 4 compounds of 7 items each one. This item grouping procedure was applied to the data obtained in the two experimental conditions (honest and faking) of the different samples analyzed (total, within- and between-subject sample). [Table 1](#) shows the compounds created.

Table 1. Compounds of the Quasi-ipsative FC Inventory

Factor	Compound	Items
ES	C1	1es-7es
	C2	8es-14es
	C3	15es-21es
	C4	22es-28es
EX	C5	1ex-7ex
	C6	8ex-14ex
	C7	15ex-21ex
	C8	22ex-28ex
OE	C9	1oe-7oe
	C10	8oe-14oe
	C11	15oe-21oe
	C12	22oe-28oe
A	C13	1a-7a
	C14	8a-14a
	C15	15a-21a
	C16	22a-28a
C	C17	1c-7c
	C18	8c-14c
	C19	15c-21c
	C20	22c-28c

Factor Analyses

To examine whether there is measurement invariance in the structure of the quasi-ipsative FC inventory, a series of non-restrictive factor analyses (that is, EFA) was carried out using the responses of participants under honest and faking conditions. Non-restrictive factor analysis refers to EFA (exploratory factor analysis) following recommendations by [Ferrando and Anguiano-Carrasco \(2010\)](#), who argue that this name is more appropriate for defining the model that is being tested in this type of factor analysis. In this sense, as they recommend, a non-restrictive factor analysis was carried out but with a confirmatory purpose, since two hypotheses are being tested.

Non-restrictive factor analyses were implemented using maximum likelihood as a model fitting procedure. To carry out these analyses, the Factor program by [Lorenzo-Seva and Ferrando \(2018\)](#) was used because it allows us to obtain absolute indices of model fit (such as root mean square of residuals, RMSR), relative model fit indices (for instance, root mean square error of approximation, RMSEA; non-normed fit index, NNFI; or comparative fit index, CFI) and to test the

Table 2. Rotated Factor Loadings of Non-restrictive Factor Analyses for the Honest and Faking Conditions Using the Total Sample

	Honest Condition (n = 939)					Faking Condition (n = 692)				
	F1	F2	F3	F4	F5	F1	F2	F3	F4	F5
C1	-.080	-.191	.520	-.139	-.084	-.174	-.065	-.061	.426	-.025
C2	.180	-.107	.559	-.045	-.213	.034	.036	.107	.509	.160
C3	-.086	.021	.519	-.066	-.265	-.104	-.237	-.031	.445	-.004
C4	-.203	.010	.610	-.086	.012	-.004	.242	-.182	.454	.045
C5	.727	-.148	-.095	-.104	-.137	-.070	-.189	.539	-.068	-.173
C6	.680	-.148	-.026	-.063	-.272	-.099	-.245	.584	.030	-.073
C7	.506	.026	-.097	-.032	-.191	-.039	-.018	.477	.031	.204
C8	.712	-.093	.007	-.142	-.134	-.118	-.152	.579	-.006	.011
C9	-.116	-.140	-.063	.702	-.197	.605	.095	-.007	.005	.048
C10	-.091	-.126	.003	.753	-.053	.642	.054	-.039	.060	-.076
C11	.002	-.096	-.035	.771	.040	.681	.161	-.005	.020	-.016
C12	-.050	.091	-.063	.774	-.103	.630	-.247	-.126	-.148	.031
C13	-.196	.653	-.021	-.123	-.213	-.274	-.601	-.202	-.285	.013
C14	-.116	.658	-.192	-.121	-.039	-.197	-.404	-.229	-.277	.192
C15	-.107	.378	.238	-.023	-.051	-.002	.072	-.047	.088	.605
C16	-.065	.631	.044	-.084	-.061	-.137	-.317	-.158	-.180	.272
C17	-.407	-.103	-.123	-.129	.401	-.091	.479	-.216	-.030	-.063
C18	-.332	-.136	-.140	-.168	.537	-.019	.661	-.205	-.009	.010
C19	-.207	-.097	-.101	-.130	.798	-.013	.702	-.194	-.089	.066
C20	-.190	-.059	-.009	-.047	.615	-.001	.627	-.190	-.017	.028

equivalence of the model. Pearson correlation matrix was used, five-factors were retained and Varimax rotation was applied to obtain orthogonal factors, in line with the theoretical framework of the Five-Factor model.

Factor analyses were also replicated using the principal component analysis (PCA) method, due to the fact that some authors, such as [Costa and McCrae \(1992\)](#) and [Goldberg \(1992\)](#), argue that PCA may be more appropriate than non-restrictive factor analyses for establishing the structure of personality inventories. The Factor program by [Lorenzo-Seva and Ferrando \(2018\)](#) was also used to perform these analyses, five factors were retained and Varimax rotation was applied.

Moreover, with the aim of verifying data fit to the proposed factor structure of the Five-Factor model, restrictive factor analyses (i.e., confirmatory factor analyses, CFA) were conducted. In this case, Lisrel 8 program ([Jöreskog & Sörbom, 1998](#)) was used, applying the maximum likelihood method. Since results were virtually the same in all cases, we have shown only the non-restrictive factor analyses solution.

Structure Congruence Analyses

In addition, following the recommendations by [Ferrando and Lorenzo-Seva \(2014\)](#), factor congruence coefficients were calculated in order to assess to what extent rotation adjustment was congruent with the target matrix proposed. That is, an analysis of factor congruence was conducted to establish the degree of similarity between factors that measure the same constructs (convergent factors) and the degree of non-similarity between factors that represent different constructs (divergent factors) in honest and faking response conditions.

Burt-Tucker's coefficients of congruence (r_c ; [Tucker, 1951](#)) were used to calculate the factor congruence, following the recommendation by [Cattell \(1971\)](#). This coefficient allows us to know the similarity of the load pattern, because if coefficients obtained between convergent factors are high (around .90) they indicate that paired factors are similar, particularly if coefficients between divergent factors are low (less than .40). For its calculation, factor loadings of rotated matrices have been used.

Results

Non-restrictive Factor Analyses of the Q15F_tri Structure in the Total Sample

[Table 2](#) shows the matrix of rotated loads in honest and faking conditions for the whole sample. With respect to the honest condition, all clusters had their highest positive loading in the factor on which they would be expected to load. Thus, the four compounds of emotional stability loaded together on factor 3, the four corresponding to extraversion loaded together on factor 1, the four compounds of openness to experience loaded together on factor 4, the four of agreeableness loaded together on factor 2, and the four compounds of conscientiousness loaded together on factor 5. Furthermore, it can be seen that none of the factors had other relevant positive loads and the negative loads were all of small magnitude. Concerning the results of non-restrictive factor analyses carried out with the responses under instructions that induced faking, the structure of five orthogonal factors was also supported by the data, although there are slight differences compared to the honest condition. Specifically, 18 of the 20 factored clusters had their highest positive loading in the factor on which they would be expected to load. The four compounds of emotional stability loaded together on factor 4, the four corresponding to extraversion loaded together on factor 3, the four openness to experience compounds loaded together on factor 1, and the four corresponding to conscientiousness loaded together on factor 2. However, only two of the agreeableness compounds (.605 for C15 and .272 for C16) had a relevant loading on the corresponding factor and the other two had a negative loading on conscientiousness. Finally, it can be observed that none of the factors had relevant additional positive loadings and that negative loadings were all of small magnitude, except for the two above mentioned agreeableness compounds.

[Table 3](#) shows the goodness of fit statistics of the model of five orthogonal factors for honest and faking conditions. Absolute fits (for example, RMSR), relative fits (e.g., RMSEA, CFI, NNFI), and equivalence statistics presented a magnitude that falls between typical good and excellent fit cut-off points for both experimental response-conditions. All estimators showed an excellent model fit in all cases,

being even slightly better for faking than honest response conditions. For example, RMSEA values were .031 and .026, NNFI values were .979 and .977, and CFI values were .989 and .988, for honest and faking conditions, respectively. Also, the size of the *t* statistic to measure the significance of RMSEA and CFI values indicated that they were statistically significant in both cases. Moreover, RMSR absolute fit indicator was .0177, 54.1% smaller than the expected value of .0327 for the honest condition, and .0209, 55% smaller than the expected value of .0380, in the case of the faking instructions.

In summary, as a whole, the results of the examination of Q15F_tri factor structure under honest and faking response-conditions for the total sample indicated that data fit a five-factor structure and that model fit indices are good or excellent in both response-conditions. Consequently, these results provide support to the invariance of the factor structure of this inventory under honest and faking conditions.

Table 3. Goodness-of-fit Indicators of the Big Five Model of Honest and Faking Conditions Using the Total Sample

	Honest	Faking
Root mean square error of approximation (RMSEA)	.031	.026
Root mean square of residuals (RMSR)	.018	.021
Non-normed fit index (NNFI)	.979	.977
Comparative fit index (CFI)	.989	.988
Goodness of fit index (GFI)	.997	.995
Adjusted goodness of fit index (AGFI)	.994	.990
Goodness of fit index without diagonal values (GFI)	.994	.986
Adjusted goodness of fit index without diagonal values (AGFI)	.988	.974

Non-restrictive Factor Analyses of the Q15F_tri Structure in the within-subject Design

The following two factor analyses presented in Table 4 were carried out using a within-subject design in order to examine if the type of experimental design can produce differences in the results of factor analyses. These analyses were carried out using a subsample of the total sample.

Results of the factor loads matrix in the honest condition replicated the factor structure shown for the total sample with notable similarity and, therefore, comments made in the case of the total sample are entirely applied to the present case. That is, all compounds had their highest positive loading on the expected factor. Thus, the five-factor structure has been supported by the data. In the case of the factor load matrix obtained under faking conditions for this design, results were very similar to those obtained for the same condition with the total sample. The five-factor structure was clearly supported by the data, although the magnitude of loads was slightly less in faking conditions than in honest conditions. Moreover, again, two of the agreeableness compounds, C13 and C14, showed reduced loadings on the factor that they would be expected to load (factor 3), with values of .034 and .267, respectively, and negative loadings on the consciousness factor (factor 2), with values of -.671 and -.483 for these clusters.

Table 5. Goodness-of-fit Indicators of the Big Five Model of Honest and Faking Conditions Using the Within-Subject Design

	Honest	Faking
Root mean square error of approximation (RMSEA)	.026	.034
Root mean square of residuals (RMSR)	.021	.025
Non-normed fit index (NNFI)	.985	.965
Comparative fit index (CFI)	.992	.982
Goodness of fit index (GFI)	.996	.993
Adjusted goodness of fit index (AGFI)	.992	.987
Goodness of fit index without diagonal values (GFI)	.991	.983
Adjusted goodness of fit index without diagonal values (AGFI)	.983	.968

Model fit indices obtained in the within-subject design are summarized in Table 5. In the honest condition, model fit values found were similar or slightly better than values found in the total sample, although differences have no practical relevance since they occurred in the thousandth values. In the faking response condition, indices were also good or excellent, although slightly lower than those of the total sample. Then, it must be taken into account that the sample size was much smaller in the present case, with results in greater sampling error.

Table 4. Rotated Factor Loadings of Non-restrictive Factor Analyses for the Honest and Faking Conditions Using the Within-subject Design (*n* = 490)

	Honest Condition					Faking Condition				
	F1	F2	F3	F4	F5	F1	F2	F3	F4	F5
C1	-.137	.485	-.112	-.123	-.175	-.074	-.032	-.031	-.169	.489
C2	-.070	.572	-.229	.173	-.115	.109	.047	.120	.070	.489
C3	-.064	.516	-.253	-.084	-.004	-.015	-.219	.040	-.087	.490
C4	-.115	.589	.028	-.206	.025	-.190	.261	.013	-.064	.418
C5	-.115	-.109	-.127	.740	-.135	.595	-.163	-.145	-.046	-.094
C6	-.082	-.012	-.285	.638	-.152	.619	-.189	-.078	-.099	.056
C7	-.008	-.085	-.173	.529	.003	.431	.003	.231	-.000	.036
C8	-.117	.010	-.143	.702	-.076	.650	-.119	.031	-.081	.001
C9	.692	-.088	-.225	-.127	-.145	-.097	.074	.031	.593	.016
C10	.773	.001	-.078	-.070	-.128	-.048	.072	-.060	.670	.054
C11	.728	-.041	.033	-.017	-.101	-.010	.171	-.018	.692	-.009
C12	.761	-.065	-.076	-.019	.115	-.071	-.280	.042	.629	-.191
C13	-.107	-.027	-.213	-.182	.629	-.149	-.671	.034	-.264	-.235
C14	-.149	-.170	-.081	-.094	.665	-.210	-.483	.267	-.196	-.226
C15	-.044	.216	-.060	-.061	.344	-.080	.091	.548	.014	.099
C16	-.047	.050	-.060	-.071	.646	-.137	-.369	.281	-.156	-.077
C17	-.148	-.146	.401	-.466	-.102	-.247	.489	-.036	-.134	-.032
C18	-.146	-.151	.546	-.323	-.154	-.215	.690	.034	-.034	-.038
C19	-.146	-.107	.805	-.212	-.088	-.244	.698	.080	-.035	-.152
C20	-.053	.005	.646	-.144	-.093	-.187	.649	.024	-.034	-.050

Table 6. Rotated Factor Loadings of Non-restrictive Factor Analyses for the Honest and Faking Conditions Using the Between-subject Design

	Honest Condition (n = 449)					Faking Condition (n = 202)				
	F1	F2	F3	F4	F5	F1	F2	F3	F4	F5
C1	.578	-.182	-.023	-.022	.096	.153	-.093	-.030	-.156	-.115
C2	.577	-.067	-.165	.204	.003	.598	.058	.128	-.045	.107
C3	.556	.089	-.231	-.052	-.048	.352	-.244	-.250	-.147	-.173
C4	.625	.017	.016	-.184	-.049	.659	.140	-.127	.144	.019
C5	-.057	-.134	-.139	.732	-.048	-.028	-.267	.331	-.157	-.191
C6	-.011	-.117	-.232	.756	-.033	-.112	-.399	.482	-.075	-.038
C7	-.083	.070	-.191	.495	-.031	.052	-.115	.584	-.116	.064
C8	.023	-.094	-.115	.728	-.127	-.043	-.209	.333	-.216	-.048
C9	-.014	-.104	-.142	-.077	.707	.006	.161	.021	.674	.042
C10	.021	-.110	-.005	-.092	.751	.019	.004	-.039	.563	-.144
C11	-.045	-.098	.029	.016	.806	.111	.110	.012	.645	-.049
C12	-.048	.086	-.119	-.047	.788	.020	-.197	-.274	.617	.047
C13	.014	.729	-.169	-.171	-.111	-.319	-.320	-.238	-.310	.141
C14	-.199	.674	.037	-.107	-.062	-.337	-.141	-.254	-.196	.141
C15	.288	.442	-.000	-.117	.026	.189	.047	.063	-.017	.576
C16	.044	.612	-.043	-.039	-.097	-.225	-.147	-.113	-.084	.462
C17	-.059	-.047	.457	-.291	-.084	-.055	.428	-.050	.034	-.144
C18	-.087	-.058	.577	-.282	-.156	.114	.524	-.192	-.019	-.079
C19	-.086	-.087	.810	-.174	-.080	.099	.647	-.009	.040	.037
C20	-.008	-.002	.602	-.206	-.030	.087	.585	-.141	.075	.011

Non-restrictive Factor Analyses of the Q15F_tri Structure in the between-Subject Design

Finally, the last two non-restrictive factor analyses, presented in Table 6 were carried out with a between-subject design, which allows us to verify the potential effect of the research design on the fit of the model, by comparing the results of these analyses with results of the within-subject design.

Again, in the honest condition, results showed that all the clusters loaded on their hypothetical factor, with relevant and positive loadings, and that they did not present positive secondary loadings on the remaining factors. It can also be seen that the remaining loadings for each factor were not significant and most of them were negative. The results under faking instructions, once again, showed that data fits a factor structure of five orthogonal factors. However, three of the compounds (C1, C13, and C14) showed a factor loading which was not significant on the factor that they would be expected to load, although they did not load significantly on any other factor.

Table 7. Goodness-of-fit Indicators of the Big Five Model of Honest and Faking Conditions Using the Between-Subject Design

	Honest	Faking
Root mean square Error of approximation (RMSEA)	.035	.023
Root mean square of residuals (RMSR)	.022	.035
Non-normed fiti (NNFI)	.973	.974
Comparative fit index (CFI)	.986	.987
Goodness of fit index (GFI)	.995	.984
Adjusted goodness of fit index (AGFI)	.990	.970
Goodness of fit index without diagonal values (GFI)	.989	.952
Adjusted goodness of fit index without diagonal values (AGFI)	.979	.908

Model fit indices for the between-subject design are presented in Table 7. Values obtained were good or excellent, with magnitudes similar to those of previous analyses. In particular, the RMSEA index should be highlighted, whose value showed a better fit under counterfeiting conditions than under honest conditions (.023 vs. .035). Therefore, in general, this latter analysis also provides em-

pirical support for the five-factor structure and for the equivalence (invariance) of this structure under honest and faking conditions.

Structure Congruence Analyses of the Q15F_tri

Table 8 presents Burt-Tucker's congruence coefficients calculated for each pair of factors in each of the two experimental conditions analysed (honest and faking). The top of the table shows Burt-Tucker's congruence coefficients for the faking condition and coefficients for the honest condition are presented at the bottom of the table. As can be seen, the results found showed very low coefficients between divergent factors in both honest and faking conditions. In all cases, values of Burt-Tucker coefficients between divergent factors were equal to or less than .40. Therefore, based on these results, factor structure remains stable in the FC inventory despite the effects of faking.

Table 8. Burt-Tucker's Congruence Coefficients among Personality Factors for each Response-Condition

	ES	EX	OE	A	C
ES	1.00	.00	-.02	.03	.09
EX	-.01	1.00	-.07	-.10	-.23
OP	-.07	-.06	1.00	-.04	.13
A	-.04	-.17	-.12	1.00	-.01
C	-.18	-.40	-.14	-.18	1.00

Note. n honest sample = 939; n faking sample = 692. The values under honest conditions appear below the diagonal and the values under faking conditions appear above the diagonal; ES = emotional stability; EX = extraversion; OE = openness to experience; A = agreeableness; C = conscientiousness.

Finally, the stability of factor congruence between convergent factors of different experimental conditions (honest vs. faking) was also analyzed. Table 9 reports the results obtained. Burt-Tucker's coefficients show much higher values between convergent factors of different experimental conditions than between divergent factors. On average, the coefficient of congruence of convergent components was .90, while average of divergent factors was .14 in absolute terms and -.10 in relative terms. In summary, these results demonstrate that the quasi-ipsative FC inventory factor structure is stable and robust against the effects of faking.

Table 9. Burt-Tucker's Congruence Coefficients among Personality Factors between the Honest and the Faking Conditions

	ES _H	EX _H	OE _H	A _H	C _H
ES _F	.96	.01	-.07	-.24	-.18
EX _F	.00	.96	-.08	-.28	-.40
OE _F	-.04	-.06	.98	-.24	-.04
A _F	.21	-.02	-.01	.73	.10
C _F	.03	-.28	-.01	-.48	.89

Note. *n* honest sample = 939; *n* faking sample = 692; ES = emotional stability; EX = extraversion; OE = openness to experience; A = agreeableness; C = conscientiousness.

Discussion

The invariance or equivalence of the measure establishes that measures in two or more groups or in two or more conditions are comparable (Millsap, 2011; Jiang et al., 2017). In the current study, the analysis of measure equivalence was conducted in order to determine if the the Five Factor measure evaluated with a quasi-ipsative FC inventory changes depending on whether the personality inventory is completed under honest or faking instructions. If there is not measurement equivalence, this would indicate that response conditions can alter construct measurement and, consequently, the instrument could not measure the same construct under honest and faking conditions. In this sense, it is pertinent in the field of Work and Organizational (W/O) Psychology to analyze this issue, particularly in the case of personality measures used in personnel selection processes since there is the possibility that applicants might voluntarily distort their answers to show a more positive self-image.

This research has contributed to the personality and faking literature in several ways. First, this is the first study that has examined measurement invariance of a FC inventory that produces algebraically non-dependent quasi-ipsative scores under faking conditions. In addition, this study is relevant because it examined measure invariance in three sub-samples and with two experimental designs (within-subject and between-subject design), which allowed us to determine the potential moderating effect of the study design on the results. So, this is the second contribution of this study.

The third contribution of this research is that examination was not limited to performing a single factor analysis; rather, three types of latent structure analyses were carried out. Non-restrictive factor analyses (maximum likelihood method), principal component analyses, and restrictive factor analyses were performed. Moreover, goodness of fit indices and Burt-Tucker's congruence coefficients were calculated to obtain a more robust index of similarities between factor structures of both conditions. Therefore, it is the most exhaustive study on this issue ever carried out with quasi-ipsative FC inventories.

The hypothesis that was tested proposed that faking behavior would not affect the factor structure of the quasi-ipsative FC inventory (Hypothesis 1). Likewise, since previous research has shown that the magnitude of faking can be affected by the type of experimental design (Martínez, 2019; Viswesvaran & Ones, 1999), a more specific analysis of the effects of faking on factor structure comparing a within-subject design and a between-subjects design was proposed (Hypothesis 2).

The findings support the robustness of the stability of the five-factor structure in the quasi-ipsative FC personality questionnaire under faking conditions. The results of the non-restrictive factor analyses showed in all cases a factor structure of five components. Fit indices obtained in these analyses showed the good fit of the model to the data. Likewise, the results of confirmatory analyses showed acceptable fit indices, both in the honest condition and in the faking condition of the two samples. These findings allow us to affirm that data fitted the proposed Five-Factor model and, therefore, factor structure remains stable even in faking conditions. This is also a unique contribution of the current study.

Furthermore, Burt-Tucker's congruence coefficients have been reported in order to provide a more robust indicator of similarities between factor structures in both experimental conditions. The results leave no room for doubt, as the values obtained showed the robustness of factor structure in FC quasi-ipsative inventories in both honest and faking conditions, even in within-subject experimental designs. When a quasi-ipsative FC inventory is used, factor structure shows convergency between honest and the faking response-conditions. Hence, this is the fifth contribution of this study.

Finally, it should also be considered that analyses were carried out under conditions that favor discrepancy in factor structures since item clusters were created through an arbitrary assignment. The use of proper compounds Big Five facets measured by the Q15F_tri would have undoubtedly contributed to a greater equivalence (invariance) of the measure and better adjustment statistics, because measurement errors and the variance of specific factors would have been reduced. However, in this research Big Five facets were not used to establish a potentially higher disagreement between honest and faking conditions. Therefore, it is another important contribution of this study, because it shows the robust structure of this inventory in one of the worst case scenarios.

Theoretical and Practical Implications

The findings of the current study have relevant implications for research and professional practice. From a theoretical point of view, this is the first study that provides empirical evidence of the effects of faking on construct validity of a quasi-ipsative FC that provides non-algebraically dependent scores. The results obtained suggest that this type of FC questionnaire is a robust instrument that controls faking effects on factor structure. Hence, these findings provide evidence of measure invariance in a quasi-ipsative FC questionnaire without algebraic dependence in faking conditions.

This study has also examined the moderating effect of the study design on the magnitude of the effects of faking on construct validity. Results shows that the consequence of faking on factor structure are equal regardless of the experimental design. Therefore, these findings, suggest that in the case of the effects of faking on the validity construct of a quasi-ipsative FC inventory the type of design does not act as a moderator.

The findings reported in this study have also implications from a practical perspective. The use of quasi-ipsative FC personality measures (without algebraic dependence) are recommended in applied contexts since the robustness of the stability of the five-factor structure in this quasi-ipsative FC inventory under faking response-instructions was supported. Therefore, using a quasi-ipsative FC questionnaire without algebraic dependence instead of an SS personality measure for high-risk decisions would allow for greater control of faking.

Limitations of the Study and Future Research

It should be noted that this study has also some limitations. First, this study was carried out in an experimental laboratory context. In this sense, it would be useful to replicate this research with other types of samples and in other contexts (e.g., candidates vs. employees) in order to explore if results vary and to obtain greater support and evidence for present results.

Second, this study has been performed with a quasi-ipsative FC personality measure that provides scores without algebraic dependence. Hence, results cannot be generalized to other types of quasi-ipsative FC inventories, that is, results obtained could be different for other quasi-ipsative formats if they provide algebraically dependent scores. Although it could be speculated, based on these findings, that the quasi-ipsative FC format as a whole

is resistant to modifications of the structure that faking produces in SS formats, it would be worthwhile to examine whether these results would be replicated with quasi-ipsative FC measures with algebraic dependence in order to provide evidence of the effects of faking on other quasi-ipsative FC inventories. Likewise, it would be useful to analyze this issue in the other FC inventory types, normative and ipsative, so as to examine whether results obtained would vary depending on FC psychometric characteristics.

Conclusions

The current study has contributed to the knowledge of the effects of faking on FC personality measures providing evidence of robustness against the effects of faking on construct validity of a quasi-ipsative FC inventory without algebraic dependence. Specifically, this study shows (1) the quasi-ipsative FC algebraically non-dependent present invariance of the measure between honest and faking response-conditions, (2) that factor congruence between convergent factors in both response-conditions is very high, obtaining high model fit indices in all cases, and (3) that the type of experimental design is not a moderator of the effects of faking on construct validity. These findings have practical relevance for the assessment of personality, especially in the area of personnel selection, where hiring decisions are more frequently affected by faking, supporting the use of quasi-ipsative FC questionnaires without algebraic dependence in order to control the effects of faking.

Finally, we encourage future researchers to replicate this study and expand these contributions using samples from different countries and other FC personality measures.

Conflict of Interest

The authors of this article declare no conflict of interest.

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